



GENE FLOW ACCROSS A GEOGRAPHICAL BARRIER

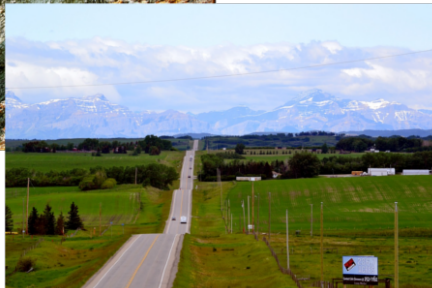
Raphaël Forien
CMAP - École Polytechnique

Les Probabilités de demain
IHES - 11 mai 2017

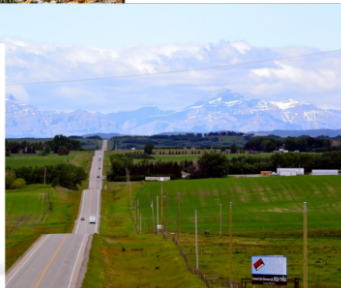
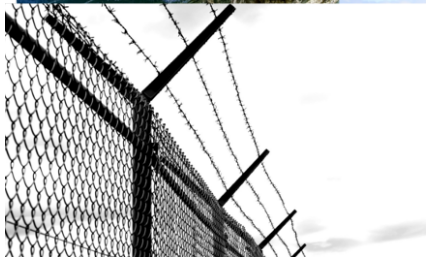
GEOGRAPHICAL BARRIERS TO DISPERSAL



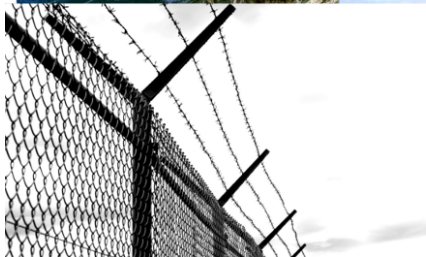
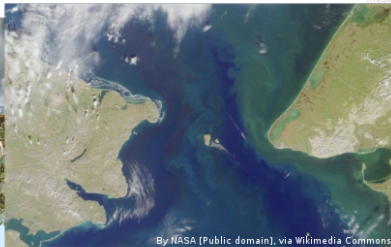
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STEPPING STONE MODEL OF DISPERSAL

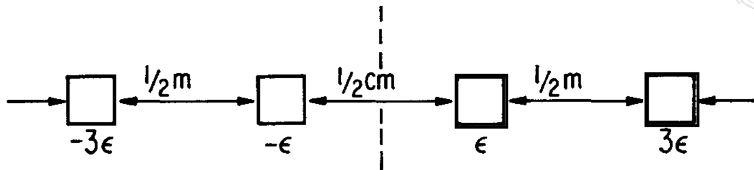


Figure: Stepping stone model with a barrier

from Nagylaki 1976

At each generation,

- the N individuals in each colony are replaced by new individuals
- a proportion $1 - m$ of them are the offspring of (uniformly chosen) parents in the same colony,
- a proportion m are the offspring of parents in neighbouring colonies

STEPPING STONE MODEL OF DISPERSAL

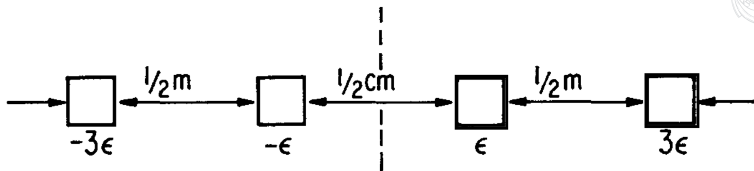


Figure: Stepping stone model with a barrier

from Nagylaki 1976

At each generation,

- the N individuals in $\pm\epsilon$ are replaced by new individuals
- a proportion $1 - \frac{1 \pm c}{2} m$ of them are the offspring of (uniformly chosen) parents in the same colony,
- a proportion $\frac{1}{2} cm$ are the offspring of parents in $\mp\epsilon$ and a proportion $\frac{1}{2} m$ come from colony $\pm 3\epsilon$

EVOLUTION OF ALLELE FREQUENCIES



Individuals are of two types, 0 and 1. Parental type is inherited by the offspring.

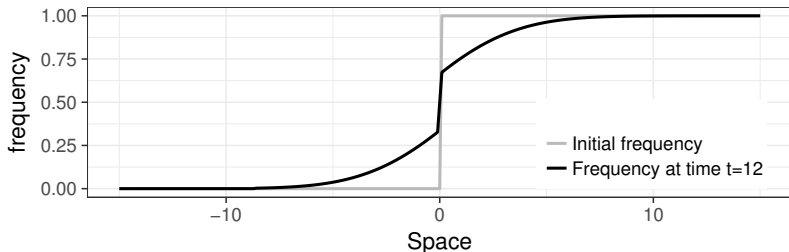


Figure: Evolution of allele frequencies with a barrier

ξ_t : position of the ancestor of a (uniformly) sampled individual t generations in the past = random walk on \mathbb{Z} with transition probabilities given by the migration matrix of the stepping stone model.

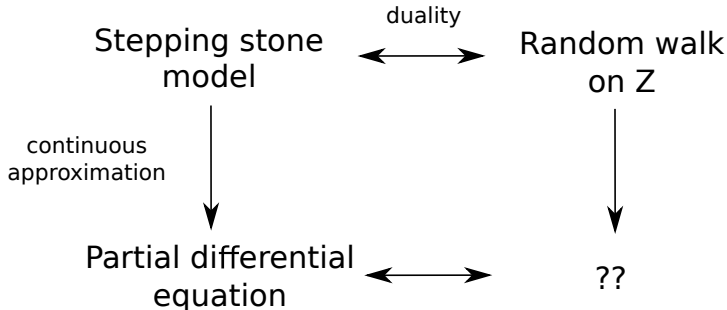
$$p(t, x) = \mathbb{P}_x (\xi_t \in [0, \infty))$$

APPLICATIONS



Goal : detect barriers to gene flow using genetic data by estimating the age of the most recent common ancestor for different pairs of individuals.

ξ_t : random walk not convenient, no explicit formulas for the law of ξ_t .



MAIN RESULT



For a sequence $(c_n)_{n \in \mathbb{N}}$, let $(\xi_n(t))_{t \geq 0}$ be a random walk on \mathbb{Z} with the corresponding transition probabilities. Set $X_n(t) = \frac{1}{\sqrt{n}} \xi_n(nt)$

Theorem 1

Suppose $\sqrt{nc_n} \xrightarrow{n \rightarrow \infty} 2\gamma \in [0, +\infty]$, then

$$X_n \xrightarrow[n \rightarrow \infty]{sko} X.$$

The process $(X(t))_{t \geq 0}$ is (the projection on \mathbb{R} of) a Markov process on $(-\infty, 0^-] \cup [0^+, +\infty)$.

When $\gamma \in (0, \infty)$, we call X partially reflected Brownian motion.

CONSTRUCTION OF PARTIALLY REFLECTED BM 1/2



Start from standard Brownian motion and keep only the excursions outside of $[-\frac{1}{2\gamma}, \frac{1}{2\gamma}]$.

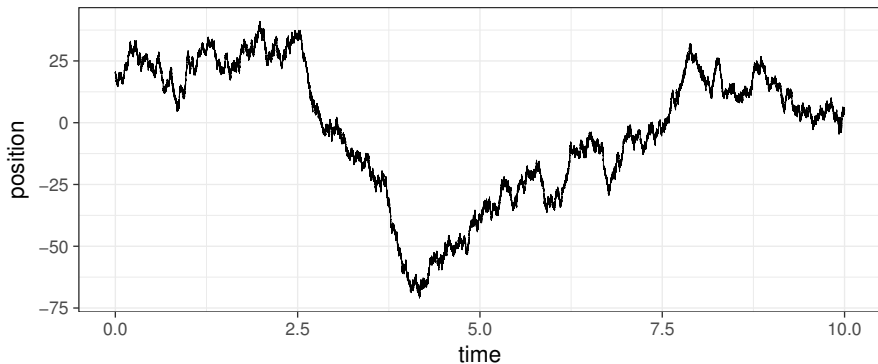


Figure: Speed and scale construction of partially reflected Brownian motion

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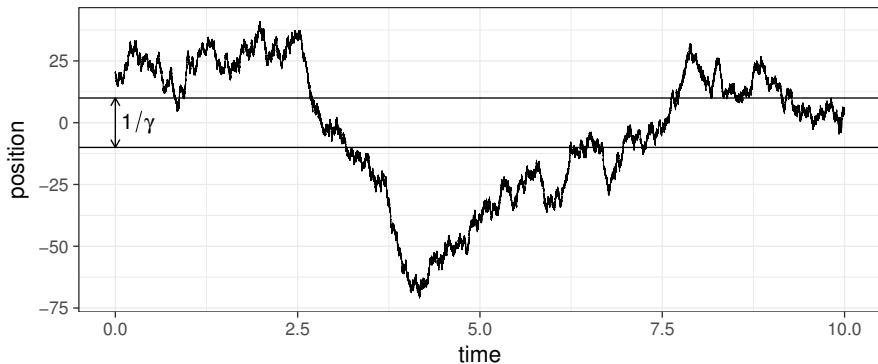


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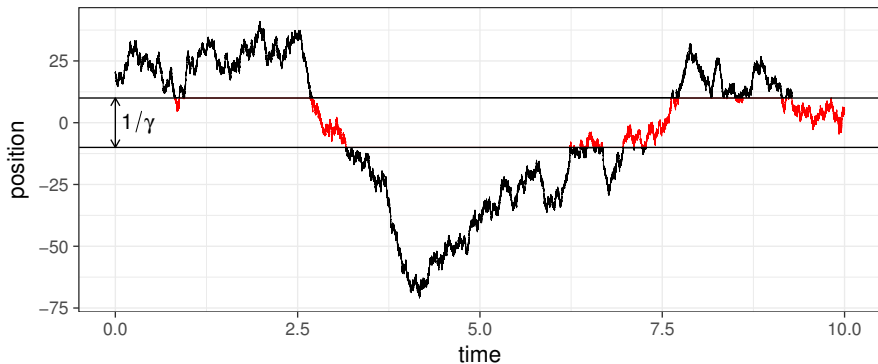


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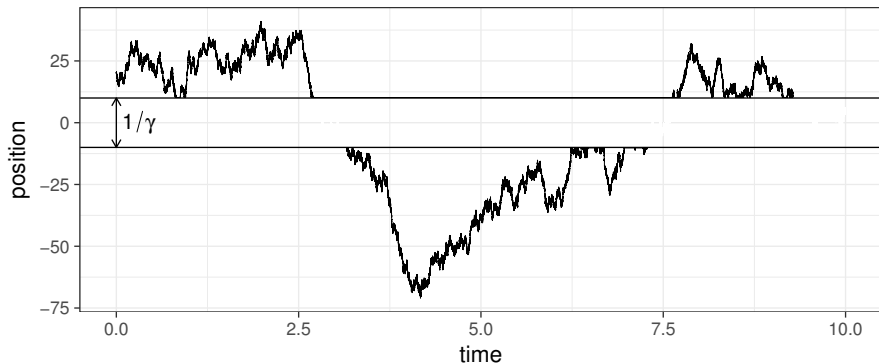


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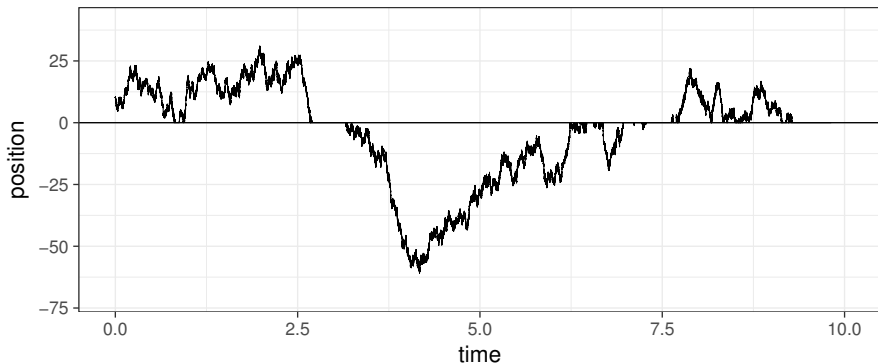


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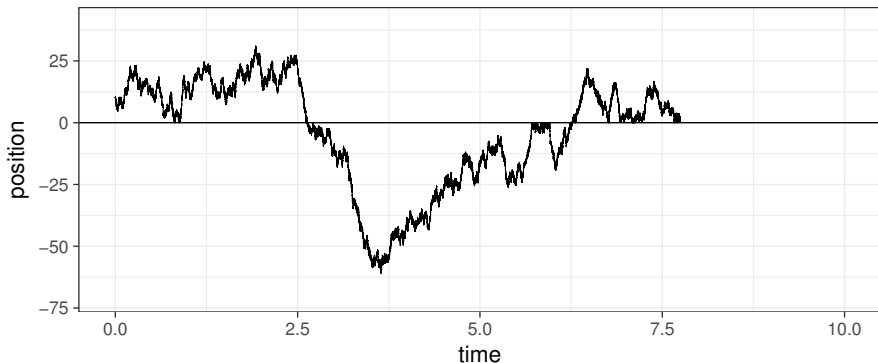
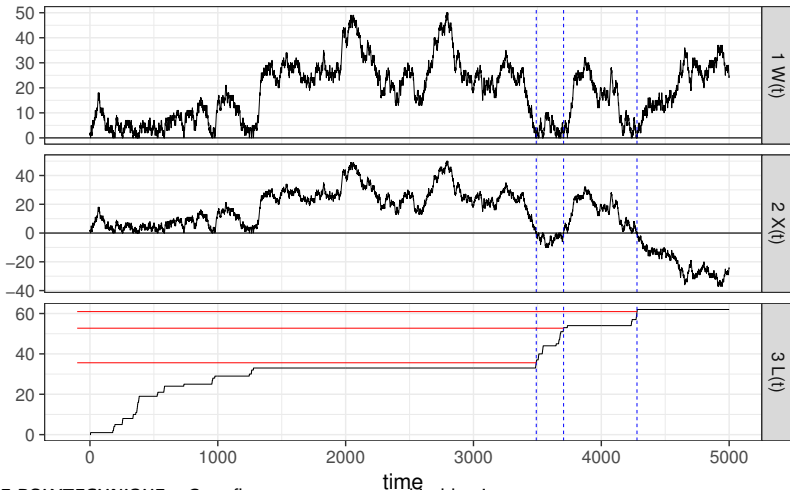


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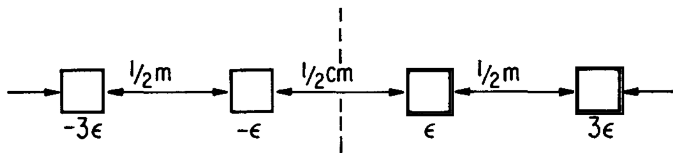
CONSTRUCTION OF PARTIALLY REFLECTED BM 2/2

Start from reflected Brownian motion $(W_t)_{t \geq 0}$, and flip it when its local time at 0 reaches an exponential variable.



SKETCH OF PROOF FOR THE CONVERGENCE RESULT

ξ_t random walk on \mathbb{Z} with transition probabilities



$$X_n(t) = \frac{1}{\sqrt{n}} \xi(nt)$$

$L_n(t) : \frac{1}{\sqrt{n}} \times$ number of visits of X_n to $\{\pm \frac{1}{\sqrt{n}}\}$ up to time t

T_i^n : time of the i -th crossing of $\{\pm \frac{1}{\sqrt{n}}\}$ by X_n

Proof of Theorem 1.

1. $|X_n|$ converges to reflected Brownian motion as $n \rightarrow \infty$,
2. $\{L_n(T_{i+1}^n) - L_n(T_i^n), i \geq 0\}$ converges to an iid sequence of $\mathcal{E}(2\gamma)$,
3. the two are asymptotically independent.

TRANSITION DENSITIES



We have an explicit formula for the transition densities of $(X_t)_{t \geq 0}$.

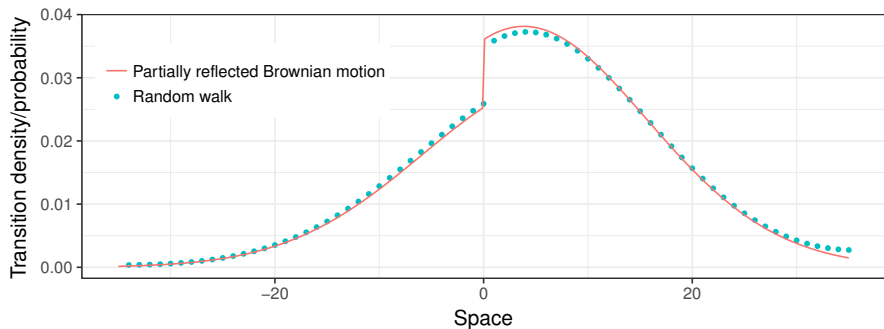
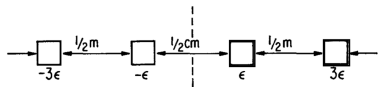


Figure: Comparison of transition probabilities for the random walk and transition densities for partially reflected Brownian motion

SUMMING UP



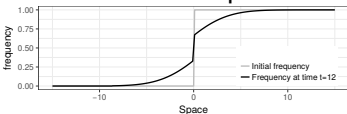
Stepping stone
model

duality
↔

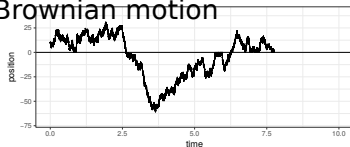
Random walk
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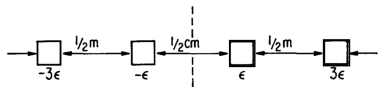
Partial differential
equation



Partially reflected
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SUMMING UP



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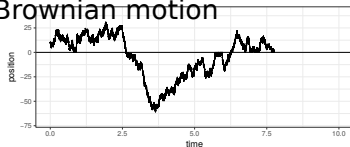
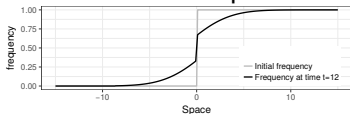
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Thank you for your attention !